

random variables

discreteprobability **mass** function

p.m.f.

$$p(x) = P(X = x)$$

$$\forall x \quad 0 \leq p(x) \leq 1$$

$$\sum_{\text{all } x} p(x) = 1$$

continuousprobability **density** function

p.d.f.

$$f(x)$$

$$\forall x \quad f(x) \geq 0$$

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

cumulative distribution function

c.d.f.

$$F(x) = P(X \leq x)$$

$$F(x) = \sum_{y \leq x} p(y)$$

$$F(x) = \int_{-\infty}^x f(y) dy$$

expected value

$$E(X) = \mu_X$$

discretecontinuous

$$E(X) = \sum_{\text{all } x} x \cdot p(x)$$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

discretecontinuous

$$E(g(X)) = \sum_{\text{all } x} g(x) \cdot p(x)$$

$$E(g(X)) = \int_{-\infty}^{\infty} g(x) \cdot f(x) dx$$

variance

$$\text{Var}(X) = \sigma_X^2 = E([X - \mu_X]^2) = E(X^2) - [E(X)]^2$$

discrete

$$\text{Var}(X) = \sum_{\text{all } x} (x - \mu_X)^2 \cdot p(x)$$

$$= \sum_{\text{all } x} x^2 \cdot p(x) - [E(X)]^2$$

continuous

$$\text{Var}(X) = \int_{-\infty}^{\infty} (x - \mu_X)^2 \cdot f(x) dx$$

$$= \left[\int_{-\infty}^{\infty} x^2 \cdot f(x) dx \right] - [E(X)]^2$$

moment-generating function

$$M_X(t) = E(e^{tX})$$

discrete

$$M_X(t) = \sum_{\text{all } x} e^{tx} \cdot p(x)$$

continuous

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} \cdot f(x) dx$$

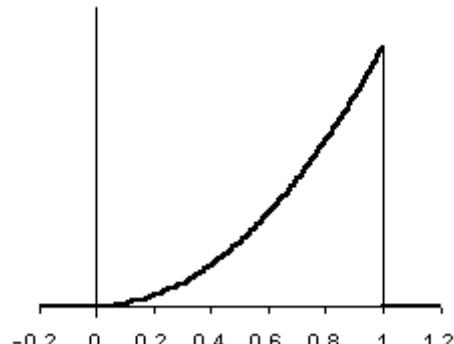
Example 1:

Let X be a continuous random variable with the probability density function

$$f(x) = k \cdot x^2, \quad 0 < x < 1,$$

$$f(x) = 0, \quad \text{otherwise.}$$

- a) What must the value of k be so that $f(x)$ is a probability density function?



1) $f(x) \geq 0,$

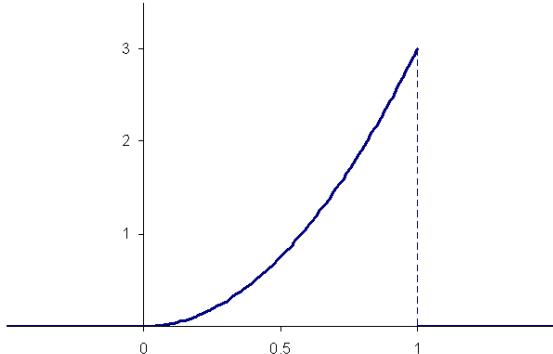
2) $\int_{-\infty}^{\infty} f(x) dx = 1.$

$$\begin{aligned}
1 &= \int_{-\infty}^{\infty} f(x) dx = \int_0^1 k \cdot x^2 dx = k \cdot \int_0^1 x^2 dx \\
&= k \cdot \left(\frac{x^3}{3} \right) \Big|_0^1 = k \cdot \left(\frac{1}{3} \right) = \frac{k}{3}. \quad \Rightarrow \quad k = 3.
\end{aligned}$$

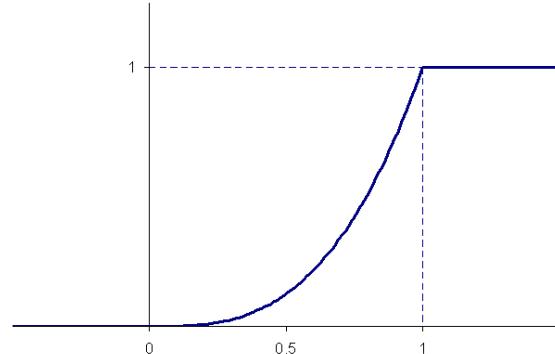
b) Find the cumulative distribution function $F(x) = P(X \leq x)$.

$$\begin{aligned}
f_X(x) &= \begin{cases} 3x^2 & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases} & x < 0 & F_X(x) = 0. \\
&& 0 \leq x < 1 & F_X(x) = \int_0^x 3y^2 dy = x^3. \\
&& x \geq 1 & F_X(x) = 1.
\end{aligned}$$

$$f_X(x)$$

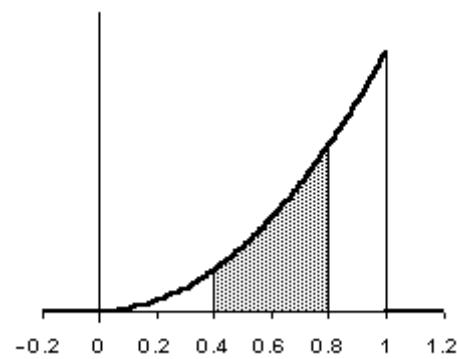


$$F_X(x)$$



c) Find the probability $P(0.4 \leq X \leq 0.8)$.

$$\begin{aligned}
P(0.4 \leq X \leq 0.8) &= \int_{0.4}^{0.8} f(x) dx \\
&= \int_{0.4}^{0.8} 3 \cdot x^2 dx = x^3 \Big|_{0.4}^{0.8} \\
&= 0.8^3 - 0.4^3 = \mathbf{0.448}.
\end{aligned}$$



OR

$$P(0.4 \leq X \leq 0.8) = F_X(0.8) - F_X(0.4) = 0.8^3 - 0.4^3 = \mathbf{0.448}.$$

- d) Find the median of the distribution of X.

Need $m = ?$ such that (Area to the left of m) $= \int_{-\infty}^m f(x)dx = \frac{1}{2}$.

$$\frac{1}{2} = \int_{-\infty}^m f(x)dx = \int_0^m 3 \cdot x^2 dx = x^3 \Big|_0^m = m^3.$$

$$m = \sqrt[3]{\frac{1}{2}} = \mathbf{0.7937}.$$

- e) Find $\mu_X = E(X)$.

$$E(X) = \mu_X = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_0^1 x \cdot (3 \cdot x^2) dx = 3 \cdot \int_0^1 x^3 dx.$$

$$= 3 \cdot \left(\frac{x^4}{4} \right) \Big|_0^1 = \frac{3}{4} = \mathbf{0.75}.$$

- f) Find $\sigma_X = SD(X)$.

$$\begin{aligned} \text{Var}(X) = \sigma_X^2 &= \left[\int_{-\infty}^{\infty} x^2 \cdot f(x) dx \right] - (\mu_X)^2 = \left[\int_0^1 3 \cdot x^4 dx \right] - \left(\frac{3}{4} \right)^2 \\ &= 3 \cdot \left(\frac{x^5}{5} \right) \Big|_0^1 - \left(\frac{3}{4} \right)^2 = \frac{3}{5} - \frac{9}{16} = \frac{3}{80} = 0.0375. \end{aligned}$$

$$\sigma_X = SD(X) = \sqrt{\text{Var}(X)} = \sqrt{0.0375} = \mathbf{0.19365}.$$

g) Find the moment-generating function of X, $M_X(t)$.

$$M_X(t) = E(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} \cdot f(x) dx = \int_0^1 e^{tx} \cdot 3x^2 dx.$$

$$u = 3x^2, \quad dv = e^{tx} dx,$$

$$du = 6x dx, \quad v = \frac{1}{t} e^{tx}.$$

$$M_X(t) = \int_0^1 e^{tx} \cdot 3x^2 dx = \left(3x^2 \cdot \frac{1}{t} e^{tx} \right) \Big|_0^1 - \int_0^1 \left(\frac{1}{t} e^{tx} \cdot 6x \right) dx$$

$$= \frac{3}{t} e^t - \int_0^1 \left(\frac{1}{t} e^{tx} \cdot 6x \right) dx$$

$$u = 6x, \quad dv = \frac{1}{t} e^{tx} dx,$$

$$du = 6 dx, \quad v = \frac{1}{t^2} e^{tx}.$$

$$\begin{aligned} M_X(t) &= \frac{3}{t} e^t - \int_0^1 \left(\frac{1}{t} e^{tx} \cdot 6x \right) dx = \frac{3}{t} e^t - \left(6x \cdot \frac{1}{t^2} e^{tx} \right) \Big|_0^1 - \int_0^1 \left(\frac{1}{t^2} e^{tx} \cdot 6 \right) dx \\ &= \frac{3}{t} e^t - \frac{6}{t^2} e^t + \left(\frac{6}{t^3} e^{tx} \right) \Big|_0^1 = \frac{3}{t} e^t - \frac{6}{t^2} e^t + \frac{6}{t^3} e^t - \frac{6}{t^3}, \quad t \neq 0. \end{aligned}$$

$$M_X(0) = 1.$$

h) Find $E(\sqrt{X})$ and $E(\ln X)$.

$$E(\sqrt{X}) = \int_0^1 \sqrt{x} \cdot 3x^2 dx = \frac{6}{7}.$$

$$E(\ln X) = \int_0^1 \ln x \cdot 3x^2 dx = -\frac{1}{3}.$$

Example 2:

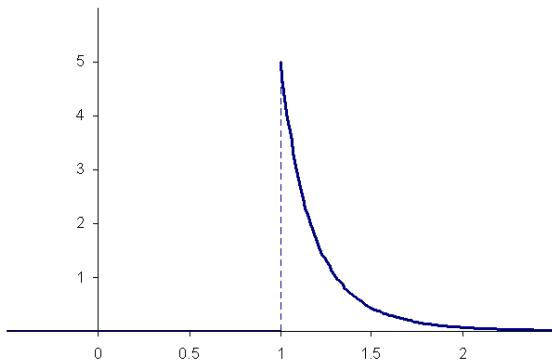
$$f_X(x) = \begin{cases} 5x^{-6} & x > 1 \\ 0 & \text{o.w.} \end{cases}$$

$x < 1 \quad F_X(x) = 0.$

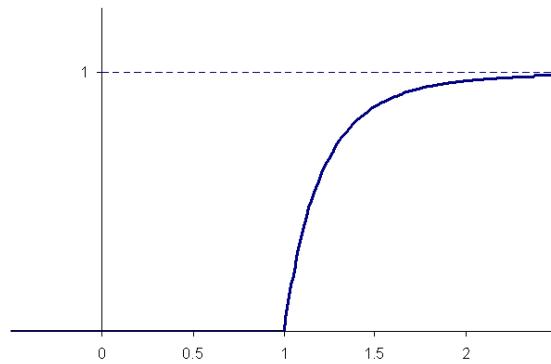
$x \geq 1 \quad F_X(x) = \int_1^x 5y^{-6} dy$

$$= -y^{-5} \Big|_1^x = 1 - x^{-5}.$$

$f_X(x)$



$F_X(x)$



$$E(X) = \mu_X = \int_1^\infty x \cdot 5x^{-6} dx = \int_1^\infty 5x^{-5} dx = \frac{5}{4} = 1.25.$$

$$E(X^2) = \int_1^\infty x^2 \cdot 5x^{-6} dx = \int_1^\infty 5x^{-4} dx = \frac{5}{3}.$$

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{5}{3} - \left(\frac{5}{4}\right)^2 = \frac{5}{48}.$$

$E(X^{10})$ does NOT exist since $\int_1^\infty x^{10} \cdot 5x^{-6} dx$ diverges.

$$\text{Median: } F_X(m) = \frac{1}{2}. \quad 1 - m^{-5} = \frac{1}{2}. \quad m = \sqrt[5]{2} \approx 1.1487.$$

$$\text{30th percentile: } F_X(\pi_{0.30}) = 0.30. \quad 1 - (\pi_{0.30})^{-5} = 0.30.$$

$$\pi_{0.30} = \sqrt[5]{\frac{1}{0.70}} \approx 1.07394.$$

Example 3:

Suppose a random variable X has the following probability density function:

$$f(x) = \begin{cases} C \cdot e^{-x} & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

- a) What must the value of C be so that $f(x)$ is a probability density function?

For $f(x)$ to be a probability density function, we must have:

$$1) \quad f(x) \geq 0, \quad 2) \quad \int_{-\infty}^{\infty} f(x) dx = 1.$$

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} f(x) dx = \int_0^1 C \cdot e^{-x} dx = C \cdot \int_0^1 e^{-x} dx \\ &= C \cdot \left[-e^{-x} \right] \Big|_0^1 = C \cdot (1 - e^{-1}) = C \cdot \left(\frac{e-1}{e} \right). \end{aligned}$$

Therefore, $C = \left(\frac{e}{e-1} \right) \approx 1.5819767.$

$$f(x) = \begin{cases} \left(\frac{e}{e-1} \right) \cdot e^{-x} & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

- b) Find $\mu_X = E(X).$

$$\mu_X = E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_0^1 x \cdot \left(\frac{e}{e-1} \right) \cdot e^{-x} dx = \left(\frac{e}{e-1} \right) \cdot \int_0^1 x \cdot e^{-x} dx.$$

Integrating by parts,

$$\begin{aligned} \int_0^1 x \cdot e^{-x} dx &= \left[(x) \cdot (-e^{-x}) \right] \Big|_0^1 - \int_0^1 (-e^{-x}) dx \\ &= -e^{-1} + \int_0^1 e^{-x} dx = -e^{-1} + \left(-e^{-x} \right) \Big|_0^1 = 1 - 2 \cdot e^{-1} = \frac{e-2}{e}. \end{aligned}$$

Therefore,

$$\mu_X = E(X) = \left(\frac{e}{e-1}\right) \cdot \int_0^1 x \cdot e^{-x} dx = \left(\frac{e}{e-1}\right) \cdot \left(\frac{e-2}{e}\right) = \frac{e-2}{e-1} \approx \mathbf{0.418}.$$

- c) Find the cumulative distribution function $F(x) = P(X \leq x)$.

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(y) dy.$$

$$F(x) = 0 \text{ for } x < 0.$$

$$F(x) = 1 \text{ for } x > 1.$$

For $0 \leq x \leq 1$,

$$\begin{aligned} F(x) &= \int_{-\infty}^x f(y) dy = \int_0^x \left(\frac{e}{e-1}\right) \cdot e^{-y} dy = \left(\frac{e}{e-1}\right) \cdot (-e^{-y}) \Big|_0^x \\ &= \left(\frac{e}{e-1}\right) \cdot (1 - e^{-x}) \\ &= \left(\frac{e}{e-1}\right) \cdot \left(\frac{e^x - 1}{e^x}\right). \end{aligned}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ \left(\frac{e}{e-1}\right) \cdot \left(\frac{e^x - 1}{e^x}\right) & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$

- d) Find the median of the probability distribution of X .

Need $m = ?$ such that $P(X \leq m) = P(X \geq m) = \frac{1}{2}$.

$$\text{Thus, } \frac{1}{2} = F(m) = \left(\frac{e}{e-1}\right) \cdot \left(\frac{e^m - 1}{e^m}\right).$$

$$\Rightarrow e^m - 1 = \left(\frac{e-1}{2 \cdot e}\right) \cdot e^m. \quad \Rightarrow e^m - \left(\frac{e-1}{2 \cdot e}\right) \cdot e^m = 1.$$

$$\Rightarrow \left(\frac{e+1}{2 \cdot e}\right) \cdot e^m = 1. \quad \Rightarrow e^m = \frac{2 \cdot e}{e+1}.$$

$$\Rightarrow m = \ln\left(\frac{2 \cdot e}{e+1}\right) \approx \mathbf{0.3799}.$$

e) Find the moment-generating function of X , $M_X(t)$.

$$\begin{aligned}
 M_X(t) &= E(e^{tX}) = \int_0^1 e^{tx} \cdot \left(\frac{e}{e-1}\right) \cdot e^{-x} dx = \left(\frac{e}{e-1}\right) \cdot \int_0^1 e^{(t-1)x} dx \\
 &= \left(\frac{e}{e-1}\right) \cdot \left(\frac{1}{t-1} \cdot e^{(t-1)x}\right) \Big|_0^1 = \left(\frac{e}{e-1}\right) \cdot \frac{1}{t-1} \cdot (e^{t-1} - 1) \\
 &= \frac{e^t - e}{(e-1) \cdot (t-1)}, \quad t \neq 1.
 \end{aligned}$$

$$M_X(1) = \frac{e}{e-1}.$$

f) Find $E(2^X)$.

$$E(2^X) = E(e^{\ln 2 \cdot X}) = M_X(\ln 2) = \frac{e^{\ln 2} - e}{(e-1) \cdot (\ln 2 - 1)} = \frac{e - 2}{(e-1) \cdot (1 - \ln 2)}.$$

Example 4:

A simple model for describing mortality in the general population in a particular country is given by the probability density function

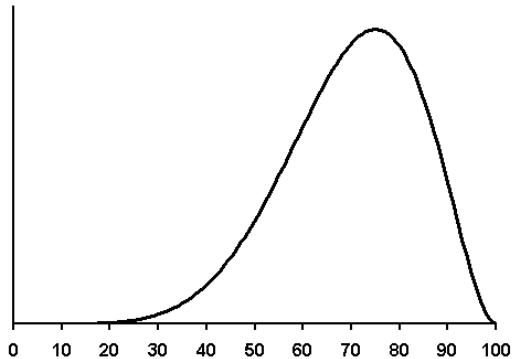
$$f(y) = \frac{252}{10^{18}} y^6 (100-y)^2, \quad 0 < y < 100.$$

- a) Verify that $f(y)$ is a valid probability density function.

1. $f(y) \geq 0$ for each y ; ✓

2. $\int_{-\infty}^{\infty} f(y) dy = 1.$

$$\begin{aligned} \int_{-\infty}^{\infty} f(y) dy &= \int_0^{100} \frac{252}{10^{18}} y^6 (100-y)^2 dy = \int_0^1 252 x^6 (1-x)^2 dx \\ &= 252 \cdot \left[\frac{1}{7} x^7 - 2 \cdot \frac{1}{8} x^8 + \frac{1}{9} x^9 \right] \Big|_0^1 = 252 \cdot \frac{2}{504} = 1. \quad \checkmark \end{aligned}$$



- b) Based on this model, which event is more likely

- or A: a person dies between the ages of 70 and 80
 B: a person lives past age 80?

$$\begin{aligned} A: \quad & \int_{70}^{80} \frac{252}{10^{18}} y^6 (100-y)^2 dy = \int_{0.7}^{0.8} 252 x^6 (1-x)^2 dx \\ &= 252 \cdot \left[\frac{1}{7} x^7 - 2 \cdot \frac{1}{8} x^8 + \frac{1}{9} x^9 \right] \Big|_{0.7}^{0.8} \end{aligned}$$

$$\approx 0.7382 - 0.4628 = 0.2754.$$

$$\begin{aligned}
 \text{B: } & \int_{80}^{100} \frac{252}{10^{18}} y^6 (100-y)^2 dy = \int_{0.8}^{1.0} 252 x^6 (1-x)^2 dx \\
 & = 252 \cdot \left[\frac{1}{7} x^7 - 2 \cdot \frac{1}{8} x^8 + \frac{1}{9} x^9 \right] \Big|_{0.8}^{1.0} \\
 & \approx 1 - 0.7382 = 0.2618.
 \end{aligned}$$

A is more likely.

- c) Given that a randomly selected individual just celebrated his 60th birthday, find the probability that he will live past age 80.

$$P(\text{over 80} \mid \text{over 60}) = \frac{P(\text{over 80} \cap \text{over 60})}{P(\text{over 60})} = \frac{\int_{80}^{100} \frac{252}{10^{18}} y^6 (100-y)^2 dy}{\int_{60}^{100} \frac{252}{10^{18}} y^6 (100-y)^2 dy}$$

$$\approx \frac{1 - 0.7382}{1 - 0.2318} = \frac{0.2618}{0.7682} \approx \mathbf{0.3408}.$$

- d) Find the value of y that maximizes $f(y)$ (**mode**).

$$\begin{aligned}
 f'(y) &= \frac{252}{10^{18}} \left[6y^5 (100-y)^2 - 2y^6 (100-y) \right] \\
 &= \frac{252}{10^{18}} y^5 (100-y) [6(100-y) - 2y] \\
 &= \frac{252}{10^{18}} y^5 (100-y) [600 - 8y] = 0.
 \end{aligned}$$

$$\Rightarrow y = 0, y = 100 \text{ (not max)}, y = \mathbf{75} \text{ years (max)}.$$

e) Find the (average) life expectancy.

$$\begin{aligned} E(Y) &= \int_{-\infty}^{\infty} y \cdot f(y) dy = \int_0^{100} \frac{252}{10^{18}} y^7 (100-y)^2 dy = \int_0^1 252 \cdot 100 x^7 (1-x)^2 dx \\ &= 252 \cdot 100 \cdot \left[\frac{1}{8} x^8 - 2 \cdot \frac{1}{9} x^9 + \frac{1}{10} x^{10} \right] \Big|_0^1 = 252 \cdot 100 \cdot \frac{2}{720} = \mathbf{70} \text{ years.} \end{aligned}$$

OR

Consider $X = \frac{Y}{100}$. Then $Y = 100X$, and X has the probability density function

$$f(x) = 252 x^6 (1-x)^2, \quad 0 < x < 1.$$

Then X has Beta distribution with $\alpha = 7$ and $\beta = 3$.

$$E(X) = \frac{\alpha}{\alpha+\beta} = \frac{7}{7+3} = 0.70. \quad E(Y) = 100 E(X) = \mathbf{70} \text{ years.}$$

f) Find the standard deviation of the lifetimes.

$$\text{Var}(X) = \frac{7 \cdot 3}{11 \cdot 10^2} = \frac{21}{1100}. \quad \text{Var}(Y) = 100^2 \text{Var}(X) = \frac{2100}{11}.$$

$$\text{SD}(Y) \approx 13.817.$$

OR

$$E(Y^2) = \int_{-\infty}^{\infty} y^2 \cdot f(y) dy = \int_0^{100} \frac{252}{10^{18}} y^8 (100-y)^2 dy = \dots$$

$$\text{Var}(Y) = E(Y^2) - [E(Y)]^2 = \dots$$

$$\text{SD}(Y) = \sqrt{\text{Var}(Y)} = \dots$$

Example 5:

Let Y denote a random variable with probability density function given by

$$f(y) = \frac{1}{2} e^{-|y|}, \quad -\infty < y < \infty. \quad (\text{double exponential p.d.f.})$$

- a) Find the moment-generating function of Y . For which values of t does it exist?

$$\begin{aligned} M_Y(t) &= \int_{-\infty}^{\infty} e^{ty} \cdot \frac{1}{2} e^{-|y|} dy = \int_{-\infty}^0 e^{ty} \cdot \frac{1}{2} e^{-|y|} dy + \int_0^{\infty} e^{ty} \cdot \frac{1}{2} e^{-|y|} dy \\ &= \int_{-\infty}^0 e^{ty} \cdot \frac{1}{2} e^y dy + \int_0^{\infty} e^{ty} \cdot \frac{1}{2} e^{-y} dy \\ &= \frac{1}{2} \int_{-\infty}^0 e^{y(t+1)} dy + \frac{1}{2} \int_0^{\infty} e^{y(t-1)} dy \end{aligned}$$

Note that the first integral converges only if $t+1 > 0$,

and the second integral converges only if $t-1 < 0$.

Therefore, the moment-generating function is only defined for $-1 < t < 1$.

$$\begin{aligned} M_Y(t) &= \frac{1}{2(t+1)} e^{y(t+1)} \Big|_{-\infty}^0 + \frac{1}{2(t-1)} e^{y(t-1)} \Big|_0^{\infty} = \frac{1}{2(t+1)} - \frac{1}{2(t-1)} \\ &= \frac{(t-1)-(t+1)}{2(t+1)(t-1)} = \frac{-2}{2(t^2-1)} = \frac{1}{1-t^2}, \quad -1 < t < 1. \end{aligned}$$

- b) Find $E(Y)$.

$$M'_Y(t) = -(1-t^2)^{-2}(-2t) = 2t(1-t^2)^{-2}$$

$$\Rightarrow E(Y) = M'_Y(0) = 0.$$

OR

$$E(Y) = \int_{-\infty}^{\infty} y \cdot \frac{1}{2} e^{-|y|} dy = 0, \quad \text{since } y \cdot \frac{1}{2} e^{-|y|} \text{ is an odd function.}$$

c) Find $\text{Var}(Y)$.

$$M_Y''(t) = 2(1-t^2)^{-2} + 2t(-2)(1-t^2)^{-3}(-2t) = \frac{2+6t^2}{(1-t^2)^3}$$

$$\Rightarrow E(Y^2) = M_Y''(0) = 2. \quad \Rightarrow \quad \text{Var}(Y) = 2 - 0^2 = 2.$$

OR

$$\text{Var}(Y) = E(Y^2) - [E(Y)]^2 = \int_{-\infty}^{\infty} y^2 \cdot \frac{1}{2} e^{-|y|} dy = \dots = 2.$$

d) Find the cumulative distribution function $F(y) = P(Y \leq y)$.

$$\text{If } y < 0, \quad F(y) = \int_{-\infty}^y \frac{1}{2} e^{-|x|} dx = \int_{-\infty}^y \frac{1}{2} e^x dx = \frac{1}{2} e^y.$$

$$\begin{aligned} \text{If } y \geq 0, \quad F(y) &= \int_{-\infty}^y \frac{1}{2} e^{-|x|} dx = \int_{-\infty}^0 \frac{1}{2} e^x dx + \int_0^y \frac{1}{2} e^{-x} dx \\ &= \frac{1}{2} + \frac{1}{2} \left(1 - e^{-y} \right) = 1 - \frac{1}{2} e^{-y}. \end{aligned}$$

Therefore,

$$F(y) = \begin{cases} \frac{1}{2} e^y & y < 0 \\ 1 - \frac{1}{2} e^{-y} & y \geq 0 \end{cases}$$

e) Find $E(Y^k)$ for positive integer k .

$$E(Y^k) = \int_{-\infty}^{\infty} y^k \cdot \frac{1}{2} e^{-|y|} dy = \int_{-\infty}^0 y^k \cdot \frac{1}{2} e^y dy + \int_0^{\infty} y^k \cdot \frac{1}{2} e^{-y} dy = \dots$$

$$\begin{matrix} k \text{ odd} \\ k \text{ even} \end{matrix} \dots = 0.$$

$$\begin{matrix} k \text{ even} \\ k \text{ odd} \end{matrix} \dots = 2 \cdot \int_0^{\infty} y^k \cdot \frac{1}{2} e^{-y} dy = \int_0^{\infty} y^k \cdot e^{-y} dy = \Gamma(k+1) = k!.$$

OR

Taylor Formula:

$$M_Y(t) = \sum_{k=0}^{\infty} \frac{t^k}{k!} M_Y^{(k)}(0) = \sum_{k=0}^{\infty} \frac{t^k}{k!} E(Y^k).$$

On the other hand,

$$M_Y(t) = \frac{1}{1-t^2} = \sum_{n=0}^{\infty} t^{2n}.$$

$$\Rightarrow \quad \text{If } k \text{ odd,} \quad E(Y^k) = 0.$$

$$\text{If } k \text{ even,} \quad k = 2n, \quad E(Y^k) = k!.$$