

STAT 400 Homework 09

Spring 2018 / Dalpiaz / UIUC

Due: Friday, April 6, 2:00 PM

Exercise 1

Let X_1, X_2, \dots, X_n be a random sample of size n from a distribution with probability density function

$$f(x, \theta) = \frac{1}{\theta} e^{-x/\theta}, \quad x > 0, \theta > 0$$

Note that, the moments of this distribution are given by

$$E[X^k] = \int_0^{\infty} \frac{x^k}{\theta} e^{-x/\theta} = k! \cdot \theta^k.$$

This will be a useful fact for Exercises 2 and 3.

(a) Obtain the maximum likelihood *estimator* of θ , $\hat{\theta}$. (This should be a function of the unobserved x_i and the sample size n .) Calculate the *estimate* when

$$x_1 = 0.50, \quad x_2 = 1.50, \quad x_3 = 4.00, \quad x_4 = 3.00.$$

(This should be a single number, for this dataset.)

Solution:

We first obtain the likelihood by **multiplying** the probability density function for each X_i . We then **simplify** this expression.

$$L(\theta) = \prod_{i=1}^n f(x_i; \theta) = \prod_{i=1}^n \frac{1}{\theta} e^{-x_i/\theta} = \theta^{-n} \exp\left(\frac{-\sum_{i=1}^n x_i}{\theta}\right)$$

Instead of directly maximizing the likelihood, we instead maximize the **log-likelihood**.

$$\log L(\theta) = -n \log \theta - \frac{\sum_{i=1}^n x_i}{\theta}$$

To maximize this function, we take a **derivative** with respect to θ .

$$\frac{d}{d\theta} \log L(\theta) = \frac{-n}{\theta} + \frac{\sum_{i=1}^n x_i}{\theta^2}$$

We set this derivative equal to **zero**, then **solve** for θ .

$$\frac{-n}{\theta} + \frac{\sum_{i=1}^n x_i}{\theta^2} = 0$$

Solving gives our *estimator*, which we denote with a **hat**.

$$\hat{\theta} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$$

Using the given data, we obtain an *estimate*.

$$\hat{\theta} = \frac{0.50 + 1.50 + 4.00 + 3}{4} = \boxed{2.25}$$

(b) Calculate the bias of the maximum likelihood *estimator* of θ , $\hat{\theta}$. (This will be a number.)

Solution:

Note that we have an exponential distribution.

$$E[X_i] = \theta$$

$$\text{Var}[X_i] = \theta^2$$

$$\begin{aligned} \text{Bias}(\hat{\theta}) &= E[\hat{\theta}] - \theta \\ &= E\left[\frac{\sum_{i=1}^n X_i}{n}\right] - \theta \\ &= \frac{1}{n} \sum_{i=1}^n E[X_i] - \theta \\ &= \frac{1}{n} n\theta - \theta \\ &= \theta - \theta = \boxed{0} \end{aligned}$$

(c) Find the mean squared error of the maximum likelihood *estimator* of θ , $\hat{\theta}$. (This will be an expression based on the parameter θ and the sample size n . Be aware of your answer to the previous part, as well as the distribution given.)

Solution:

$$\begin{aligned} \text{MSE}(\hat{\theta}) &= [\text{Bias}(\hat{\theta})]^2 + \text{Var}(\hat{\theta}) \\ &= 0 + \text{Var}\left(\frac{\sum_{i=1}^n X_i}{n}\right) \\ &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) \\ &= \frac{1}{n^2} n\theta^2 = \boxed{\frac{\theta^2}{n}} \end{aligned}$$

(d) Provide an *estimate* for $P[X > 4]$ when

$$x_1 = 0.50, x_2 = 1.50, x_3 = 4.00, x_4 = 3.00.$$

Solution:

$$P[X > 4] = e^{-4/\theta}$$

$$\hat{P}[X > 4] = e^{-4/\hat{\theta}} = e^{-4/2.25} = \boxed{0.1690}$$

Exercise 2

Let X_1, X_2, \dots, X_n be a random sample of size n from a distribution with probability density function

$$f(x, \alpha) = \alpha^{-2} x e^{-x/\alpha}, \quad x > 0, \alpha > 0$$

(a) Obtain the maximum likelihood *estimator* of α , $\hat{\alpha}$. Calculate the *estimate* when

$$x_1 = 0.25, \quad x_2 = 0.75, \quad x_3 = 1.50, \quad x_4 = 2.5, \quad x_5 = 2.0.$$

Solution:

We first obtain the likelihood by **multiplying** the probability density function for each X_i . We then **simplify** this expression.

$$L(\alpha) = \prod_{i=1}^n f(x_i; \alpha) = \prod_{i=1}^n \alpha^{-2} x_i e^{-x_i/\alpha} = \alpha^{-2n} \left(\prod_{i=1}^n x_i \right) \exp\left(-\frac{\sum_{i=1}^n x_i}{\alpha}\right)$$

Instead of directly maximizing the likelihood, we instead maximize the **log-likelihood**.

$$\log L(\alpha) = -2n \log \alpha + \sum_{i=1}^n \log x_i - \frac{\sum_{i=1}^n x_i}{\alpha}$$

To maximize this function, we take a **derivative** with respect to α .

$$\frac{d}{d\alpha} \log L(\alpha) = \frac{-2n}{\alpha} + \frac{\sum_{i=1}^n x_i}{\alpha^2}$$

We set this derivative equal to **zero**, then **solve** for α .

$$\frac{-2n}{\alpha} + \frac{\sum_{i=1}^n x_i}{\alpha^2} = 0$$

Solving gives our *estimator*, which we denote with a **hat**.

$$\hat{\alpha} = \frac{\sum_{i=1}^n x_i}{2n} = \frac{\bar{x}}{2}$$

Using the given data, we obtain an *estimate*.

$$\hat{\alpha} = \frac{0.25 + 0.75 + 1.50 + 2.50 + 2.0}{2 \cdot 5} = \boxed{0.70}$$

(b) Obtain the method of moments *estimator* of α , $\tilde{\alpha}$. Calculate the *estimate* when

$$x_1 = 0.25, x_2 = 0.75, x_3 = 1.50, x_4 = 2.5, x_5 = 2.0.$$

Solution:

We first obtain the first **population moment**. Notice the integration is done by identifying the form of the integral is that of the second moment of an exponential distribution.

$$E[X] = \int_0^{\infty} x \cdot \alpha^{-2} x e^{-x/\alpha} dx = \frac{1}{\alpha} \int_0^{\infty} \frac{x^2}{\alpha} e^{-x/\alpha} dx = \frac{1}{\alpha} (2\alpha^2) = 2\alpha$$

We then set the first population moment, which is a function of α , equal to the first **sample moment**.

$$2\alpha = \frac{\sum_{i=1}^n x_i}{n}$$

Solving for α , we obtain the method of moments *estimator*.

$$\tilde{\alpha} = \frac{\sum_{i=1}^n x_i}{2n} = \frac{\bar{x}}{2}$$

Using the given data, we obtain an *estimate*.

$$\tilde{\alpha} = \frac{0.25 + 0.75 + 1.50 + 2.50 + 2.0}{2 \cdot 5} = \boxed{0.70}$$

Note that, in this case, the MLE and MoM estimators are the same.

Exercise 3

Let X_1, X_2, \dots, X_n be a random sample of size n from a distribution with probability density function

$$f(x, \beta) = \frac{1}{2\beta^3} x^2 e^{-x/\beta}, \quad x > 0, \beta > 0$$

(a) Obtain the maximum likelihood *estimator* of β , $\hat{\beta}$. Calculate the *estimate* when

$$x_1 = 2.00, x_2 = 4.00, x_3 = 7.50, x_4 = 3.00.$$

Solution:

We first obtain the likelihood by **multiplying** the probability density function for each X_i . We then **simplify** this expression.

$$L(\beta) = \prod_{i=1}^n f(x_i; \beta) = \prod_{i=1}^n \frac{1}{2\beta^3} x_i^2 e^{-x_i/\beta} = 2^{-n} \beta^{-3n} \left(\prod_{i=1}^n x_i \right) \exp \left(\frac{-\sum_{i=1}^n x_i}{\beta} \right)$$

Instead of directly maximizing the likelihood, we instead maximize the **log-likelihood**.

$$\log L(\beta) = -n \log 2 - 3n \log \beta + \sum_{i=1}^n \log x_i - \frac{\sum_{i=1}^n x_i}{\beta}$$

To maximize this function, we take a **derivative** with respect to β .

$$\frac{d}{d\beta} \log L(\beta) = \frac{-3n}{\beta} + \frac{\sum_{i=1}^n x_i}{\beta^2}$$

We set this derivative equal to **zero**, then **solve** for β .

$$\frac{-3n}{\beta} + \frac{\sum_{i=1}^n x_i}{\beta^2} = 0$$

Solving gives our *estimator*, which we denote with a **hat**.

$$\hat{\beta} = \frac{\sum_{i=1}^n x_i}{3n} = \frac{\bar{x}}{3}$$

Using the given data, we obtain an *estimate*.

$$\hat{\beta} = \frac{2.00 + 4.00 + 7.50 + 3.00}{3 \cdot 4} = \boxed{1.375}$$

(b) Obtain the method of moments *estimator* of β , $\tilde{\beta}$. Calculate the *estimate* when

$$x_1 = 2.00, x_2 = 4.00, x_3 = 7.50, x_4 = 3.00.$$

Solution:

We first obtain the first **population moment**. Notice the integration is done by identifying the form of the integral is that of the third moment of an exponential distribution.

$$E[X] = \int_0^{\infty} x \cdot \frac{1}{2\beta^3} x^2 e^{-x/\beta} dx = \frac{1}{2\beta^2} \int_0^{\infty} \frac{x^3}{\beta} e^{-x/\beta} dx = \frac{1}{2\beta^2} (6\beta^3) = 3\beta$$

We then set the first population moment, which is a function of β , equal to the first **sample moment**.

$$3\beta = \frac{\sum_{i=1}^n x_i}{n}$$

Solving for β , we obtain the method of moments *estimator*.

$$\tilde{\beta} = \frac{\sum_{i=1}^n x_i}{3n} = \frac{\bar{x}}{3}$$

Using the given data, we obtain an *estimate*.

$$\tilde{\beta} = \frac{2.00 + 4.00 + 7.50 + 3.00}{3 \cdot 4} = \boxed{1.375}$$

Note again, the MLE and MoM estimators are the same.

Exercise 4

Let X_1, X_2, \dots, X_n be a random sample of size n from a distribution with probability density function

$$f(x, \lambda) = \lambda x^{\lambda-1}, \quad 0 < x < 1, \lambda > 0$$

(a) Obtain the maximum likelihood *estimator* of λ , $\hat{\lambda}$. Calculate the *estimate* when

$$x_1 = 0.10, \quad x_2 = 0.20, \quad x_3 = 0.30, \quad x_4 = 0.40.$$

Solution:

We first obtain the likelihood by **multiplying** the probability density function for each X_i . We then **simplify** this expression.

$$L(\lambda) = \prod_{i=1}^n f(x_i; \lambda) = \prod_{i=1}^n \lambda x_i^{\lambda-1} = \lambda^n \left(\prod_{i=1}^n x_i \right)^{\lambda-1}$$

Instead of directly maximizing the likelihood, we instead maximize the **log-likelihood**.

$$\log L(\lambda) = n \log \lambda + (\lambda - 1) \sum_{i=1}^n \log x_i$$

To maximize this function, we take a **derivative** with respect to λ .

$$\frac{d}{d\lambda} \log L(\lambda) = \frac{n}{\lambda} + \sum_{i=1}^n \log x_i$$

We set this derivative equal to **zero**, then **solve** for β .

$$\frac{n}{\lambda} + \sum_{i=1}^n \log x_i = 0$$

Solving gives our *estimator*, which we denote with a **hat**.

$$\hat{\lambda} = -\frac{n}{\sum_{i=1}^n \log x_i}$$

Using the given data, we obtain an *estimate*.

$$\hat{\lambda} = -\frac{n}{\sum_{i=1}^n \log x_i} = -\frac{4}{\log(0.1 \cdot 0.2 \cdot 0.3 \cdot 0.4)} = \boxed{0.6631}$$

Note that this is actually a reparameterization of an example seen in class where $\lambda = \frac{1}{\theta}$. Had you realized this, you could have simply found the answer via invariance.

(b) Obtain the method of moments *estimator* of λ , $\tilde{\lambda}$. Calculate the *estimate* when

$$x_1 = 0.10, \quad x_2 = 0.20, \quad x_3 = 0.30, \quad x_4 = 0.40.$$

Solution:

We first obtain the first **population moment**.

$$E[X] = \int_0^1 x \cdot \lambda x^{\lambda-1} dx = \frac{\lambda}{\lambda+1}$$

We then set the first population moment, which is a function of β , equal to the first **sample moment**.

$$\frac{\lambda}{\lambda+1} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$$

Solving for λ , we obtain the method of moments *estimator*.

$$\tilde{\lambda} = \frac{\bar{x}}{1 - \bar{x}}$$

Using the given data, we obtain an *estimate*.

$$\bar{x} = \frac{0.1 + 0.2 + 0.3 + 0.4}{4} = 0.25$$

$$\tilde{\lambda} = \frac{0.25}{1 - 0.25} = \boxed{\frac{1}{3}}$$

Note that the MLE and MoM *estimators* are different.